

考試科目	數理統計	所別	統計所	考試時間	6月29日 上午第一節 星期
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(20%) 1. Let X_1, X_2, X_3 be independent with X_i having density $f(x_i) = \exp(-x_i)$, $x_i > 0$.

Let $U_1 = X_1 + X_2 + X_3$, $U_2 = X_2/X_1$, and $U_3 = X_3/U_1$.

(a) Find the joint density of (U_1, U_2, U_3) .

(b) Discuss the independency of U_1, U_2 , and U_3 .

(30%) 2. Suppose that $X_i \sim N(i\theta, \sigma^2)$ are independent, $i=1, 2, \dots, n$. Let $U = \frac{\sum_{i=1}^n iX_i}{k}$, where $k = \sum_{i=1}^n i^2$, and $V^2 = \frac{\sum_{i=1}^n (X_i - iU)^2}{(n-1)}$.

(a) Show that $U \sim N(\theta, \sigma^2/k)$

(b) Show that $(n-1)V^2/\sigma^2 \sim \chi_{n-1}^2$.

(c) Show that U and V^2 are independent.

(20%) 3. Suppose that a sample of size n is drawn from the uniform distribution

$$f(x) = \begin{cases} \frac{1}{\theta}, & 0 < x < \theta, \\ 0, & \text{otherwise.} \end{cases}$$

(a) Show that $X_{(n)} = \max(X_1, X_2, \dots, X_n)$ is a biased estimator for θ .

(b) Find a constant c , so that $cX_{(n)}$ would be an unbiased estimator for θ .

(30%) 4. Let X_1, \dots, X_n be independent, with $X_i \sim T(d, \beta)$, where d and β are both unknown. (i.e. the p.d.f. of X_i is $f(x_i) = \frac{x_i^{d-1} \exp(-x_i/\beta)}{\beta^d \Gamma(d)}$, $x_i > 0$, $d > 0$, $\beta > 0$).

(a) Find the best unbiased estimator of $d\beta$.

(b) Find the best unbiased estimator of $d^d \beta^d$.

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考試科目	線性模式	所別	統計系	考試時間	6月24日 上午第 節 星期二
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考慮線性模式 $\begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{pmatrix} = Y = X\beta + \epsilon$, 其中 X 是秩 (rank)

為 p 的 $n \times p$ 矩陣, $E[\epsilon] = 0$, $D[\epsilon] = \sigma^2 I_n$. 令 $\hat{\beta}$ 表示 β 的 LSE, $\hat{Y} = X\hat{\beta} = \begin{pmatrix} \hat{Y}_1 \\ \vdots \\ \hat{Y}_n \end{pmatrix}$, $e = Y - \hat{Y}$.

(1) (i) 求 $D[e]$ (5%) (ii) 證明: $\sum_{i=1}^n \hat{Y}_i (Y_i - \hat{Y}_i) = 0$ (5%)

(iii) 證明: $E(\|e\|^2) = (n-p)\sigma^2$, 其中 $\|e\|^2 = e'e$. (15%)

(敘述出所要引用的定理)

接著, 設 A 是已知的 $g \times p$ 矩陣, 且 A 的秩為 p , c 是已知的 $g \times 1$ 向量.

(2) (i) 在 $A\beta = c$ 的限制下, 推導出 β 的 LSE $\hat{\beta}_H$. (20%)

(ii) 證明 $\text{Var}(\hat{\beta}_{Hi}) \leq \text{Var}(\hat{\beta}_i)$, 其中 $\hat{\beta}_{Hi}$ 與 $\hat{\beta}_i$ 分別是 $\hat{\beta}_H$ 與 $\hat{\beta}$ 的第 i 個分量. (15%)

(iii) 令 $RSS = \|Y - X\hat{\beta}\|^2$, $RSS_H = \|Y - X\hat{\beta}_H\|^2$

證明: $RSS_H - RSS = (A\hat{\beta} - c)' [A(X'X)^{-1}A']^{-1} (A\hat{\beta} - c)$ (15%)

現進一步設 $\epsilon \sim N_n(0, \sigma^2 I_n)$.

(3) 證明: 在假設 $H: A\beta = c$ 成立時,

$$\frac{(RSS_H - RSS)/g}{RSS/(n-p)} \sim F_{g, n-p}. \quad (25\%)$$

(敘述出所要引用的定理)

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