

考試科目	數理統計	所別	統計所	考試時間	6月23日 上午第一節 星期二
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1. Suppose X_1, X_2, \dots are i.i.d. r.v.'s from $B(1, p)$. ($>0\%$)
- (a) Show that there does not exist an unbiased estimator of p^{-1} , given a fixed sample, say, X_1, \dots, X_n .
- (b) Let N be the first n such that $X_n = 1$. Show that $E(N) = p^{-1}$.
2. Let $\{X_n\}, n = 1, 2, \dots$ be a sequence of random variables and X be a random variable. ($>0\%$)
- (a) Define converges in probability ($X_n \xrightarrow{P} X$), converges in distribution ($X_n \xrightarrow{d} X$), and converges in quadratic mean ($X_n \xrightarrow{q.m.} X$).
- (b) Show that $X_n \xrightarrow{P} X$ implies $X_n \xrightarrow{d} X$, and $X_n \xrightarrow{q.m.} X$ implies $X_n \xrightarrow{d} X$.
- (c) Find counter examples that neither converges in probability implies converges in quadratic mean, nor converges in quadratic mean implies converges in probability.

3. Let X be a r.v. having the geometric distribution; that is, ($>0\%$)

$$f(x; \theta) = \theta(1 - \theta)^x, \quad x = 0, 1, \dots, \quad \theta \in \Theta = (0, 1),$$

and let $U(X)$ be defined as follows: $U(X) = 1$ if $X = 0$ and $U(X) = 0$ if $X \neq 0$. Show that $U(X)$ is a UMVU estimator of θ and conclude that it is an unreasonable one.

4. Let X_1, \dots, X_n be independent r.v.'s with p.d.f. given by ($>0\%$)

$$f(x; \theta) = \frac{1}{\theta} e^{-x/\theta} I_{(0, \infty)}, \quad \theta \in \Theta = (0, \infty).$$

- (a) Derive the UMP test for testing the hypothesis $H_0: \theta \geq \theta_0$ against the alternative $H_1: \theta < \theta_0$ at the level of significance α .
- (b) Determine the minimum sample size n required to obtain power at least 0.95 against the alternative $\theta_1 = 500$ when $\theta_0 = 1,000$ and $\alpha = 0.05$.
5. Let X_1, \dots, X_n be i.i.d. r.v.'s from $U(0, \theta)$ and set $R = X_{(n)} - X_{(1)}$, where ($>0\%$)

$$X_{(n)} = \max_{1 \leq i \leq n} X_i \quad \text{and} \quad X_{(1)} = \min_{1 \leq i \leq n} X_i.$$

- (a) Find the distribution of R .
- (b) Show that a confidence interval of θ , based on R , with confidence coefficient $1 - \alpha$ is of the form $[R, R/c]$, where c is a root of the equation

$$c^{n-1}[n - (n-1)c] = \alpha.$$

考試科目	線性模式	所別	統計	考試時間	6月23日 上午 第二節
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以線性模式 $y = \beta_0 + \beta_1 x_1 + \dots + \beta_{p-1} x_{p-1} + \epsilon$ 配合適
 n 個資料點 $\{(x_{i1}, x_{i2}, \dots, x_{i,p-1}; y_i) \mid i=1, 2, \dots, n\}$, 即
 資料點滿足 $y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_{p-1} x_{i,p-1} + \epsilon_i, i=1, 2, \dots, n$,
 或 $y = X\beta + \epsilon \dots \dots (1)$

其中 $y = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}, X = \begin{pmatrix} 1 & x_{11} & \dots & x_{1,p-1} \\ 1 & x_{21} & \dots & x_{2,p-1} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_{n1} & \dots & x_{n,p-1} \end{pmatrix}, \epsilon = \begin{pmatrix} \epsilon_1 \\ \vdots \\ \epsilon_n \end{pmatrix},$

$n > p$, $\text{rank}(X) = p$, 且 $E(\epsilon) = 0, \text{Cov}(\epsilon) = \sigma^2 I_n$.

10%一. 令 e_i 表示第 i 個殘差值, $i=1, 2, \dots, n$, 證明 $\sum_{i=1}^n e_i = 0$.

10%二. 令 $SS_E = e'e$, 其中 $e' = (e_1, \dots, e_n)$, 證明 $E(SS_E) = (n-p)\sigma^2$, 並敘述你所引用的定理.

10%三. 定義 $R^2 = 1 - \frac{SS_E}{\sum_{i=1}^n (y_i - \bar{y})^2}$, 證明 $y = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}$ 與 $\hat{y} = \begin{pmatrix} \hat{y}_1 \\ \vdots \\ \hat{y}_n \end{pmatrix}$

$= Hy$ 的相關係數之平方等於 R^2 , 其中

$H = X(X'X)^{-1}X'$.

10%四. 令 $H = (h_{ij})$, 證明 $\forall 1 \leq i \leq n, 0 \leq h_{ii} \leq 1$, 且 $\sum_{i=1}^n h_{ii} = p$.

五. 設 $A: q \times p$ 是秩(rank)為 q 的已知矩陣, $c: q \times 1$ 是已知向量. (i) 在 $A\beta = c$ 的限制下, 證明 β 的 LSE

$$\tilde{\beta} = \hat{\beta} + (X'X)^{-1}A' [A(X'X)^{-1}A']^{-1} (c - A\hat{\beta})$$

10% 其中 $\hat{\beta} = (X'X)^{-1}X'y$.

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10% (ii) 在有限制條件下 與沒有限制條件下的殘差平方和, 那一個較大? 詳細證明你的理由。

10% (iii) 設 $\epsilon \sim N_n(0, \sigma^2 I_n)$. 詳細敘述在檢定假設 $H: A\beta = c$ 時, 所要用到的統計量。

六. 現在考慮增加一些迴歸自變數, 即將模式(1)擴大如下:

$$\begin{aligned} y &= X\beta + Zr + \epsilon \quad \dots \dots \dots (2) \\ &= (X, Z) \begin{pmatrix} \beta \\ r \end{pmatrix} + \epsilon \\ &= W\delta + \epsilon \end{aligned}$$

其中 $W = (X, Z)$, $\delta = \begin{pmatrix} \beta \\ r \end{pmatrix}$, 且 $W: n \times (p+t)$ 的秩為 $n+t$.

令 $\hat{\delta}_G = \begin{pmatrix} \hat{\beta}_G \\ \hat{r}_G \end{pmatrix}$ 是 δ 的 LSE.

15% (i) 證明 $\hat{\beta}_G = \hat{\beta} - L\hat{r}_G$, $\hat{r}_G = (Z'RZ)^{-1}Z'R(Y - X\hat{\beta})$, 其中

$$L = (X'X)^{-1}X'Z, \quad R = I_n - X(X'X)^{-1}X'$$

15% (ii) 模式(1) 與模式(2) 的殘差平方和 那一個較大? 詳細證明你的理由。